

A lower bound for irredundant Ramsey numbers¹

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Abstract

Given a graph $G=(V,E)$, a vertex subset $U \subseteq V$ is called *irredundant* if every vertex $v \in U$ either has no neighbours in U or there exists a vertex $w \in V \setminus U$ such that v is the only neighbour of w in U . The irredundant Ramsey number $s(m,n)$ is the smallest N such that any red–blue edge colouring of K^N yields either an m -element irredundant subset in the blue graph or an n -element irredundant subset in the red graph. Using probabilistic methods we show that

$$s(m,n) > c_m \left(\frac{n}{\log n} \right)^{(m^2-m-1)[2(m-1)]}.$$

1. Introduction

Let $G=(V,E)$ be a graph with vertex set V and edge set E . A subset of vertices $U \subseteq V$ is called *irredundant* if every vertex $v \in U$ either has no neighbours in U or there exists a vertex $w \in V \setminus U$ such that v is the only neighbour of w in U (in this case w is called a *private neighbour* of v). For every pair of integers $m,n \geq 2$ the *irredundant Ramsey number* $s(m,n)$ is the smallest integer N such that in any red–blue colouring of the edges of a complete graph K^N on N vertices either the blue graph contains an m -element irredundant subset or the red graph contains an n -element irredundant subset. Define also the *mixed Ramsey number* $t(m,n)$ as the smallest N such that any red–blue colouring of the edges of K^N yields an m -element irredundant subset in the blue graph or an n -element independent subset in the red graph. Recall that the *Ramsey number* $r(m,n)$ is the smallest N such that any red–blue colouring

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of the edges of K^N contains either a blue copy of K^m or a red copy of K^n . Since an independent set is clearly irredundant, the above definitions imply that

$$s(m, n) \leq t(m, n) \leq r(m, n) \quad (1)$$

for all admissible m, n .

Irredundant Ramsey numbers were introduced in [2]. In [3] asymptotic lower bounds for the diagonal irredundant Ramsey numbers $s(n, n)$ and the off-diagonal mixed Ramsey numbers $t(m, n)$, $m < n$, were obtained. In particular, it was shown that

$$t(m, n) > c_m \left(\frac{n}{\log n} \right)^{(m^2 - m - 1)/[2(m-1)]}, \quad (2)$$

this result was obtained by using the so-called *probabilistic method* (see [1] as a general reference). It was also shown in [3] that

$$t(3, n) \leq \frac{\sqrt{10}}{2} n^{3/2}.$$

The exact values for some small irredundant Ramsey numbers are known and can be found, e.g., in [3].

The purpose of this note is to establish an asymptotic lower bound for the off-diagonal irredundant Ramsey numbers $s(m, n)$, where m is fixed and n tends to infinity. We prove the following result.

Theorem 1. *For every $m \geq 3$ there exists a positive constant c_m such that*

$$s(m, n) > c_m \left(\frac{n}{\log n} \right)^{(m^2 - m - 1)/[2(m-1)]}.$$

This result matches (up to a constant factor) the bound (2) for the mixed Ramsey numbers $t(m, n)$ and thus implies (2) in view of (1).

Our proof is also based on the probabilistic method and applies *large deviation inequalities*. A similar approach has already been used in [7] for obtaining asymptotic lower bounds for various Ramsey-type numbers. We discuss this approach in Section 2. In Section 3 the proof of the main result is presented.

We end this section with some notation used in the sequel. We denote by $[N]$ the set $\{1, \dots, N\}$. The complete graph on $[N]$ is denoted by K^N . For every two disjoint vertex subsets $S, T \subseteq V(G)$ let $E(S)$ be the edge set of the subgraph of G spanned by S , $E(S, T)$ is the set of all edges of G between S and T , $e(S) = |E(S)|$ and $e(S, T) = |E(S, T)|$. A red–blue colouring of the edges of K^N induces the red graph $\langle R \rangle$ and the blue graph $\langle B \rangle$. We denote by $\langle U \rangle_R$ ($\langle U \rangle_B$, resp.) the induced subgraph of $\langle R \rangle$ ($\langle B \rangle$, resp.) on U .

For a fixed graph H we define

$$\rho(H) = \max_{H' \subseteq H, |H'| > 2} \frac{e(H') - 1}{|H'| - 2}$$

(in order to avoid trivialities throughout the paper we always assume that $e(H) \geq 2$). For a finite family of fixed graphs $\mathcal{H} = \{H_1, \dots, H_l\}$ the *density* of the family $\rho(\mathcal{H})$ is

$$\rho(\mathcal{H}) = \min\{\rho(H_i) : 1 \leq i \leq l\}.$$

Given a family $\mathcal{H} = \{H_1, \dots, H_l\}$, a graph G is called \mathcal{H} -free if it does not contain a copy of H_i for every $1 \leq i \leq l$.

2. Large deviation inequalities

Roughly speaking, large deviation inequalities assert that under certain conditions a random variable X is highly concentrated near its mean and its tail probabilities are exponentially small.

The simplest example of a large deviation inequality is the bound on the tail of a binomial distribution, essentially due to Chernoff [4]. If X is the sum of n mutually independent indicator random variables each taking the value 1 with probability p and the value 0 with probability $1 - p$, then the expectation of X equals np and for every constant $0 < \varepsilon < 1$ the following inequalities hold:

$$\Pr[X < (1 - \varepsilon)np] < e^{-\varepsilon^2 np/2}, \quad (3)$$

$$\Pr[X > (1 + \varepsilon)np] < e^{-\varepsilon^2 (1 + \varepsilon)np/2}. \quad (4)$$

When X is the sum of many ‘rarely dependent’ indicator random variables, it is also possible in certain cases to obtain exponential bounds on the tails of X . Let us describe a general scheme first presented in [6].

Suppose Q is a finite universal set (in our instances Q is the edge set of a complete graph on N vertices). Let $\{J_i : i \in Q\}$ be a set of independent indicator random variables, $\Pr[J_i = 1] = p_i$ for every $i \in Q$ ($J_i = 1$ if the corresponding edge belongs to $E(G)$, where G is a random graph on N vertices in which each edge is chosen independently with probability p). Let $\{Q(\alpha)\}_{\alpha \in I}$ be a family of subsets of Q , where I is a finite index set. Define $X_\alpha = \prod_{i \in Q(\alpha)} J_i$ (then $X_\alpha = 1$ if and only if all the edges of $Q(\alpha)$ belong to $E(G)$). Now define

$$X = \sum_{\alpha \in I} X_\alpha,$$

(in our instances X counts the number of subgraphs of G having some specified properties).

We shall make use of the bound on the upper tail of another random variable X_0 which is tightly connected to X and is defined as

$$X_0 = \max\{r : \exists \text{ distinct } \alpha_1, \dots, \alpha_r \in I \text{ with } X_{\alpha_i} = 1 \\ \text{and } Q(\alpha_i) \cap Q(\alpha_j) = \emptyset, i \neq j\}.$$

Clearly, $X_0 \leq X$. Let $\mu = EX$ be the expectation of X , then the following holds (see [5]):

Claim 1.

$$\Pr[X_0 \geq k] \leq \frac{\mu^k}{k!}$$

for every natural k .

For the sake of completeness we repeat the short proof.

Proof.

$$\begin{aligned} \Pr[X_0 \geq k] &\leq \sum^1 \Pr[(X_{\alpha_1} = 1) \wedge \cdots \wedge (X_{\alpha_k} = 1)] \\ &= \frac{1}{k!} \sum^2 \Pr[(X_{\alpha_1} = 1) \wedge \cdots \wedge (X_{\alpha_k} = 1)] \\ &= \frac{1}{k!} \sum^2 \Pr[X_{\alpha_1} = 1] \cdots \Pr[X_{\alpha_k} = 1] \\ &\leq \frac{1}{k!} \sum^3 \Pr[X_{\alpha_1} = 1] \cdots \Pr[X_{\alpha_k} = 1] = \frac{\mu^k}{k!}, \end{aligned}$$

where \sum^1 is over sets of k mutually independent events $X_{\alpha_i} = 1$, while \sum^2 is over ordered k -tuples of mutually independent events and \sum^3 is over all ordered k -tuples of events. \square

In particular, we deduce from the above claim that

$$\Pr[X_0 \geq 5\mu] < \left(\frac{e}{5}\right)^{5\mu}. \quad (5)$$

(It is worth noting that in certain cases one can also obtain exponential bounds on the lower tail of X_0 , see, e.g., [7]. However, the above cited simple bound will suffice for our purposes here).

3. Asymptotic lower bounds for $s(m, n)$

Recall that we are treating the off-diagonal irredundant Ramsey numbers $s(m, n)$, that is, m is fixed while n tends to infinity.

The proof of the main result is a simple consequence of the following:

Lemma 1. *Let $\mathcal{H} = \{H_1, \dots, H_l\}$ be a family of fixed graphs with density $\rho(\mathcal{H}) > 0$. Then there exists a constant $c = c(\mathcal{H})$ such that for every sufficiently large integer N there exists a graph G_0 on N vertices having the following properties:*

1. G_0 is \mathcal{H} -free;

2. G_0 has no independent set of size $n = \lceil cN^{1/\rho(\mathcal{H})} \ln N \rceil$;
3. for every two disjoint subsets of vertices $S, T \subseteq V(G_0)$ of size $|S| = |T| = n$ one has $e(S, T) > n$.

Proof. For every $1 \leq i \leq l$ let H'_i be a subgraph of H_i such that $\rho(H'_i) = \rho(H_i)$. Denoting $\mathcal{H}' = \{H'_1, \dots, H'_l\}$, note that if G_0 is \mathcal{H}' -free then it is clearly \mathcal{H} -free, therefore we may assume that $\rho(H_i) = (e(H_i) - 1) / (|H_i| - 2)$, $1 \leq i \leq l$. For every $1 \leq i \leq l$ set $v_i = |H_i|$, $e_i = e(H_i)$. Set also

$$e_{\min} = \min\{e_i : 1 \leq i \leq l\},$$

$$e_{\max} = \max\{e_i : 1 \leq i \leq l\}.$$

Consider a random graph $G(N, p)$ — a graph on N labelled vertices in which all edges are chosen independently with probability p . We set with foresight $p = c_0 N^{-1/\rho(\mathcal{H})}$, where $0 < c_0 < 1$ is a sufficiently small constant.

For every two disjoint subsets $S, T \subseteq V(G)$ of size $|S| = |T| = n$ we define the following random variables. First, let $X_S = e(S)$, $X_{S,T} = e(S, T)$. Also, denote by Y_S the number of subgraphs, each isomorphic to one of the graphs from \mathcal{H} and having at least one edge inside S , and by Z_S the maximal number of pairwise edge disjoint subgraphs, each isomorphic to one of the graphs from \mathcal{H} and having at least one edge inside S . Let $Y_{S,T}$ denote the number of subgraphs, each isomorphic to one of the graphs from \mathcal{H} and having at least one edge in $E(S, T)$, and let $Z_{S,T}$ denote the maximal number of pairwise edge disjoint subgraphs, each isomorphic to one of the graphs from \mathcal{H} and having at least one edge in $E(S, T)$. Clearly, $Y_S \geq Z_S$ and $Y_{S,T} \geq Z_{S,T}$. Denote by A_S the event $X_S > e_{\max} Z_S$ and by $A_{S,T}$ the event $X_{S,T} > e_{\max} Z_{S,T} + n$.

Claim 2. If A_S holds for every $S \subseteq V$ of size $|S| = n$ and $A_{S,T}$ holds for every pair of disjoint subsets $S, T \subseteq V(G)$ of size $|S| = |T| = n$, then G contains a subgraph G_0 on N vertices, satisfying the requirements of the lemma.

Proof. Let \mathbf{H} be a maximal (under inclusion) family of pairwise edge disjoint subgraphs of G , each isomorphic to one of the graphs from \mathcal{H} . Deleting all edges of all graphs from \mathbf{H} we clearly obtain an \mathcal{H} -free graph G_0 on N vertices. Denote by \mathbf{H}_S , $|S| = n$, the subfamily of \mathbf{H} , consisting of all subgraphs from \mathbf{H} , having at least one edge in $E(S)$, and by $\mathbf{H}_{S,T}$, $|S| = |T| = n$, the subfamily of \mathbf{H} , consisting of all subgraphs from \mathbf{H} , sharing at least one edge with $E(S, T)$. Obviously, $|\mathbf{H}_S| \leq Z_S$, $|\mathbf{H}_{S,T}| \leq Z_{S,T}$. While deleting the edges of subgraphs from \mathbf{H} , we delete at most $e_{\max} |\mathbf{H}_S| \leq e_{\max} Z_S$ edges from $E(S)$ and at most $e_{\max} |\mathbf{H}_{S,T}| \leq e_{\max} Z_{S,T}$ edges from $E(S, T)$, hence the subgraph G_0 satisfies also the conditions (2) and (3) of the lemma. \square

Now our aim is to show that under appropriate choice of the constants c_0 and c the inequality $\Pr[\bigwedge_{|S|=n} A_S \wedge \bigwedge_{|S|=|T|=n} A_{S,T}] > 0$ holds. To this end, we show that the random variables $X_S, Z_S, X_{S,T}, Z_{S,T}$ are highly concentrated around their means and

hence if, say, $EX_S > 10e_{\max}EZ_S$ and $EX_{S,T} > 10e_{\max}EZ_{S,T}$, then both probabilities $\Pr[\bar{A}_S]$ and $\Pr[\bar{A}_{S,T}]$ are exponentially small. This will imply that the probability that either there exists some set S for which \bar{A}_S holds or there exists a pair S, T for which $\bar{A}_{S,T}$ holds is less than 1.

The random variable X_S is clearly binomially distributed with parameters $\binom{n}{2}$ and p , therefore from (3) we obtain for every $0 < \varepsilon < 1$

$$\Pr\left[X_S < (1 - \varepsilon)\binom{n}{2}p\right] < e^{-\varepsilon^2\binom{n}{2}p/2}. \quad (6)$$

Similarly, $X_{S,T}$ is binomially distributed with parameters n^2 and p , and hence (3) implies that

$$\Pr[X_{S,T} < (1 - \varepsilon)n^2p] < e^{-\varepsilon^2n^2p/2}. \quad (7)$$

Now we bound the upper tail of Z_S by using Claim 1. To this end, we estimate EY_S . Note that

$$Y_S = Y_{S,1} + \dots + Y_{S,l},$$

where $Y_{S,i}$ is the number of copies of H_i , having at least one edge in $E(S)$. Representing $Y_{S,i}$ as a sum of indicator random variables we can write

$$\binom{n}{2} \binom{N-n}{v_i-2} p^{e_i} \leq EY_{S,i} \leq \binom{n}{2} \binom{N-2}{v_i-2} v_i! p^{e_i},$$

therefore (recalling that $n = o(N)$ and hence $\binom{N-n}{v_i-2} = \Theta(N^{v_i-2})$) we have

$$c_{i,1} \binom{n}{2} p^{(N^{v_i-2})/(e_i-1)} p^{e_i-1} \leq EY_{S,i} \leq c_{i,2} \binom{n}{2} p^{(N^{v_i-2})/(e_i-1)} p^{e_i-1},$$

where $c_{i,1}$ and $c_{i,2}$ are some positive constants depending only on H_i .

The definitions of $\rho(\mathcal{H})$ and p imply that

$$c_1 c_0^{e_{\max}-1} \binom{n}{2} p \leq EY_S \leq c_2 c_0^{e_{\min}-1} \binom{n}{2} p,$$

where $c_1 = c_1(\mathcal{H})$ and $c_2 = c_2(\mathcal{H})$ are positive constants.

Substituting in (5) EY_S and Z_S instead of μ, X_0 , respectively, we conclude that

$$\Pr[Z_S \geq 5EY_S] \leq e^{-5(\ln 5-1)EY_S}. \quad (8)$$

Turning to estimating the upper tail of $Z_{S,T}$ we act in a quite similar manner. Taking c_1 sufficiently small and $c_2 > 1$ sufficiently large we can write

$$c_1 c_0^{e_{\max}-1} n^2 p \leq EY_{S,T} \leq c_2 c_0^{e_{\min}-1} n^2 p,$$

Also, (5) implies that

$$\Pr[Z_{S,T} \geq 5EY_{S,T}] \leq e^{-5(\ln 5-1)EY_{S,T}}. \quad (9)$$

Comparing EX_S and EY_S , $EX_{S,T}$ and $EY_{S,T}$ we observe

$$\frac{1}{c_2 c_0^{e_{\min}-1}} \leq \frac{EX_S}{EY_S}, \frac{EX_{S,T}}{EY_{S,T}} \leq \frac{1}{c_1 c_0^{e_{\max}-1}}.$$

Let us choose c_0 so that the expression $c_2 c_0^{e_{\min}-1}$ will be equal to, say, $1/10e_{\max}$. Then

$$10e_{\max} \leq \frac{EX_S}{EY_S}, \frac{EX_{S,T}}{EY_{S,T}} \leq \frac{c_2}{c_1} c_0^{-e_{\max}+e_{\min}} 10e_{\max}. \quad (10)$$

Now, by (6) with $\varepsilon = 1/2$, (8) and (10)

$$\begin{aligned} \Pr[\bar{A}_S] &= \Pr[X_S \leq e_{\max} Z_S] \leq \Pr\left[X_S \leq \frac{EX_S}{2}\right] + \Pr\left[e_{\max} Z_S \geq \frac{EX_S}{2}\right] \\ &\leq \Pr\left[X_S \leq \frac{EX_S}{2}\right] + \Pr[Z_S \geq 5EY_S] \\ &\leq e^{-\binom{n}{2}p/8} + e^{-\frac{c_1 c_0^{-e_{\min}+e_{\max}}}{10c_2 e_{\max}} [5 \ln 5 - 5] \binom{n}{2} p} \leq 2e^{-c_3 n^2 p} \end{aligned}$$

for some constant $c_3 > 0$.

Also, (6) with $\varepsilon = \frac{1}{3}$, (9) and (10) imply for sufficiently large N (noting that $EX_{S,T}/n \rightarrow \infty$)

$$\begin{aligned} \Pr[\bar{A}_{S,T}] &= \Pr[X_{S,T} \leq e_{\max} Z_{S,T} + n] \\ &\leq \Pr\left[X_{S,T} \leq \frac{EX_{S,T}}{2} + n\right] + \Pr\left[e_{\max} Z_{S,T} \geq \frac{EX_{S,T}}{2}\right] \\ &\leq \Pr\left[X_{S,T} \leq \frac{2EX_{S,T}}{3}\right] + \Pr[Z_{S,T} \geq 5EY_{S,T}] \\ &\leq e^{-n^2 p/18} + e^{-\frac{c_1 c_0^{-e_{\min}+e_{\max}}}{10c_2 e_{\max}} [5 \ln 5 - 5] n^2 p} \leq 2e^{-c_3 n^2 p} \end{aligned}$$

(taking c_3 small enough).

Therefore,

$$\begin{aligned} \Pr[\exists S : \bar{A}_S] &\leq \binom{N}{n} 2e^{-c_3 n^2 p}, \\ \Pr[\exists S, T : \bar{A}_{S,T}] &\leq \binom{N}{n}^2 2e^{-c_3 n^2 p}. \end{aligned}$$

Using the inequality $\binom{N}{n} \leq \left(\frac{eN}{n}\right)^n$, we write

$$\binom{N}{n} 2e^{-c_3 n^2 p} < \left(\frac{N}{n}\right)^2 2e^{-c_3 n^2 p} < \left(\frac{eN}{n} 2e^{-c_3 n p/2}\right)^{2n}.$$

Taking c sufficiently large it follows that

$$\Pr\left[\bigwedge_{|S|=n} A_S \wedge \bigwedge_{|S|=|T|=n} A_{S,T}\right] > 0. \quad \square$$

Returning to the proof of the main result, we modify its formulation slightly for the sake of convenience and prove that

$$s(m, 2n - 2) > c'_m \left(\frac{n}{\log n} \right)^{(m^2 - m - 1)/[2(m - 1)]},$$

where $c'_m > 0$ is a constant depending only on m . Denote $H_0 = K^m$, $H_i = K^{m-i} + (K^{i,i} - iK^2)$, $2 \leq i \leq m$, where $G_1 + G_2$ denotes the join of G_1 and G_2 and $K^{i,i} - iK^2$ is obtained from the complete bipartite graph $K^{i,i}$ by removing a perfect matching. Denote $\mathcal{H} = \{H_0, H_2, \dots, H_m\}$. The density of \mathcal{H} is easily computable and equals to $(m^2 - m - 1)/(2(m - 1))$. Consider an \mathcal{H} -free graph G_0 on N vertices $[N]$ having the properties stated in the preceding lemma. We colour the edges of G_0 red and the edges of \bar{G}_0 blue. Now we claim that $\langle B \rangle$ does not contain an irredundant set of size m and $\langle R \rangle$ does not contain an irredundant set of size $2n - 2$. As observed in [3], if $\langle B \rangle$ contains an m -element irredundant subset U , then either U is independent in $\langle B \rangle$ (in this case $\langle U \rangle_R = K^m = H_0$) or for some $2 \leq i \leq m$ U contains a subset U_0 of size $|U_0| = i$ such that $U \setminus U_0$ is independent in $\langle B \rangle$ and every vertex v of U_0 has a private neighbour w relative to U (in this case, denoting by W_0 the set of the private neighbours of the vertices from U_0 , we can easily see that $\langle U \cup W_0 \rangle_R$ contains a copy of H_i), hence $\langle R \rangle$ contains one of the graphs from \mathcal{H} . Therefore, since G_0 is \mathcal{H} -free, $\langle B \rangle$ indeed does not contain any irredundant set of size m .

Consider now a set $U \subseteq [N]$ of size $|U| = 2n - 2$. Since $\alpha(G_0) < n$, there are at least n non-isolated vertices in $\langle U \rangle_R$. Fix a subset $U_0 \subseteq U$ of size $|U_0| = n$, whose members are non-isolated vertices in $\langle U \rangle_R$. If U is irredundant, then clearly there is a subset W_0 of size $|W_0| = n$ such that the induced bipartite graph $\langle U_0, W_0 \rangle_R$ consists of a matching of size n and thus contains exactly n edges — a contradiction with the properties of G_0 . Therefore, $\langle R \rangle$ does not contain any irredundant set of size $2n - 2$. \square

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